

The Minimax Rate of HSIC Estimation for Translation-Invariant Kernels*

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Quick Summary

- Hilbert-Schmidt independence criterion (HSIC [1, 2, 3, 4]; aka. distance covariance): popular dependency measure.
- Applications: feature selection, causal discovery, independence testing, clustering, sensitivity analysis, uncertainty quantification.
- Many known estimators converge at a rate of $\mathcal{O}_P(n^{-1/2})$.
- Contribution: For a large class of distributions and kernels on \mathbb{R}^d , $\mathcal{O}_P(n^{-1/2})$ is the optimal rate.

HSIC

- Given $X = (X_m)_{m=1}^M \sim \mathbb{P}$ on $\mathcal{X} = \times_{m=1}^M \mathcal{X}_m$, \mathcal{X}_m is equipped with kernel k_m and feature map $\phi_{k_m} : \mathcal{X}_m \rightarrow \mathcal{H}_{k_m}$, HSIC takes the form

$$\text{HSIC}_k(\mathbb{P}) = \left\| \mu_k(\mathbb{P}) - \mu_k \left(\otimes_{m=1}^M \mathbb{P}_m \right) \right\|_{\mathcal{H}_k}, \quad k := \otimes_{m=1}^M k_m$$

with $\otimes_{m=1}^M \mathbb{P}_m$ the product of the marginal distributions \mathbb{P}_m , $m \in [M] := \{1, \dots, M\}$, and $\mu_k(\mathbb{P}) = \mathbb{E}_{X \sim \mathbb{P}}[\phi_k(X)]$.

- We focus on $\mathcal{X} = \mathbb{R}^d$, $\mathcal{X}_m = \mathbb{R}^{d_m}$, $d = \sum_{m=1}^M d_m$; $\mathbb{R}^d = \times_{m=1}^M \mathbb{R}^{d_m}$.
- Gaussian kernels: $k_m(\mathbf{x}, \mathbf{y}) = e^{-\gamma \|\mathbf{x} - \mathbf{y}\|_{\mathbb{R}^{d_m}}^2}$ ($\gamma > 0$).
- Translation-invariant kernels: There exists $\psi_m : \mathbb{R}^{d_m} \rightarrow \mathbb{R}$ such that $k_m(\mathbf{x}, \mathbf{y}) = \psi_m(\mathbf{x} - \mathbf{y})$. Example: Gaussian kernel.

Our Goal: Lower Bound

- $\hat{F}_n :=$ any estimator of $\text{HSIC}_k(\mathbb{P})$ based on n i.i.d. samples from \mathbb{P} .
- A positive sequence $(\xi_n)_{n=1}^\infty$ is a lower bound of HSIC estimation if there exists $c > 0$ such that

$$\inf_{\hat{F}_n} \sup_{\mathbb{P} \in \mathcal{P}} \mathbb{P}^n \left\{ \left| \text{HSIC}_k(\mathbb{P}) - \hat{F}_n \right| \geq c \xi_n \right\} > 0 \quad \forall n.$$

worst distribution
best estimator

- An estimator with a matching upper bound is called minimax-optimal.
- Wanted: $\xi_n \asymp n^{-1/2}$.

Tool: Le Cam's Method

- Key [5]: There exist $\alpha > 0$ and a positive sequence $(s_n)_{n=1}^\infty$ such that for any fixed n , there exists an adversarial pair $(\mathbb{P}_{\theta_0}, \mathbb{P}_{\theta_1}) \in \mathcal{P} \times \mathcal{P}$ s.t.
 - KL $\left(\mathbb{P}_{\theta_1}^n \parallel \mathbb{P}_{\theta_0}^n \right) \leq \alpha$, and
 - $|\text{HSIC}_k(\mathbb{P}_{\theta_1}) - \text{HSIC}_k(\mathbb{P}_{\theta_0})| \geq 2s_n > 0$.
- Then, for all n ,

$$\inf_{\hat{F}_n} \sup_{\mathbb{P} \in \mathcal{P}} \mathbb{P}^n \left\{ \left| \text{HSIC}_k(\mathbb{P}) - \hat{F}_n \right| \geq s_n \right\} \geq \max \left(\frac{e^{-\alpha}}{4}, \frac{1 - \sqrt{\alpha/2}}{2} \right).$$

Our Adversarial Pair

- Let \mathcal{G} be $\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ Gaussians on $\mathbb{R}^d = \times_{m=1}^M \mathbb{R}^{d_m}$ with covariance

$$\boldsymbol{\Sigma} = \boldsymbol{\Sigma}(i, j, \rho) = \begin{bmatrix} 1 & \cdots & 0 & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & & \vdots \\ 0 & \cdots & 1 & \rho & \cdots & 0 \\ 0 & \cdots & \rho & 1 & \cdots & 0 \\ \vdots & & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & 0 & \cdots & 1 \end{bmatrix} \in \mathbb{R}^{d \times d},$$

where $i = d_1$, $j = d_1 + 1$, $\rho \in (-1, 1)$.

- We choose $\mathbb{P}_{\theta_0} = \mathcal{N}(\boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0)$ and $\mathbb{P}_{\theta_1} = \mathcal{N}(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1)$ with

$$\boldsymbol{\mu}_0 = \mathbf{0}_d \in \mathbb{R}^d, \quad \boldsymbol{\Sigma}_0 = \boldsymbol{\Sigma}(d_1, d_1 + 1, 0) = \mathbf{I}_d \in \mathbb{R}^{d \times d},$$

$$\boldsymbol{\mu}_1 = \frac{1}{\sqrt{dn}} \mathbf{1}_d \in \mathbb{R}^d, \quad \boldsymbol{\Sigma}_1 = \boldsymbol{\Sigma}(d_1, d_1 + 1, \rho_n) \in \mathbb{R}^{d \times d},$$

with $\rho_n = \frac{1}{\sqrt{n}}$.

Proof Sketch

- We use the reduction

$$\sup_{\mathbb{P} \in \mathcal{P}} \mathbb{P}^n \left\{ \left| \text{HSIC}_k(\mathbb{P}) - \hat{F}_n \right| \geq s_n \right\} \geq \sup_{\mathbb{P} \in \mathcal{G}} \mathbb{P}^n \left\{ \left| \text{HSIC}_k(\mathbb{P}) - \hat{F}_n \right| \geq s_n \right\}.$$

- For our adversarial pair $(\mathbb{P}_{\theta_0}, \mathbb{P}_{\theta_1})$, one can show that

- KL $\left(\mathbb{P}_{\theta_1}^n \parallel \mathbb{P}_{\theta_0}^n \right) \leq \alpha := \frac{5}{4}$ for $n \geq 2$ (Gaussians \Rightarrow closed-form), and
- $|\text{HSIC}_k(\mathbb{P}_{\theta_1}) - \text{HSIC}_k(\mathbb{P}_{\theta_0})| \geq 2s_n := \frac{2c}{\sqrt{n}} > 0$.

Main Result

- Let \mathcal{P} be any class of Borel probability measures containing the d -dimensional Gaussians, $k = \otimes_{m=1}^M k_m$ with $k_m : \mathbb{R}^{d_m} \times \mathbb{R}^{d_m} \rightarrow \mathbb{R}$ continuous bounded shift-invariant characteristic kernels. Then, there exists a constant $c > 0$, such that for any $n \geq 2$

$$\inf_{\hat{F}_n} \sup_{\mathbb{P} \in \mathcal{P}} \mathbb{P}^n \left\{ \left| \text{HSIC}_k(\mathbb{P}) - \hat{F}_n \right| \geq \frac{c}{\sqrt{n}} \right\} \geq \frac{1 - \sqrt{\frac{5}{8}}}{2}.$$

- Gaussian k_m -s: $c = \frac{\gamma}{2(2\gamma+1)^{\frac{d+1}{4}}} > 0$.
- General case by Bochner's theorem ($c > 0$).

Discussion

- Many of the existing HSIC estimators on \mathbb{R}^d are minimax-optimal.
- Existing lower bounds (MMD [6], mean embedding [7], covariance operator [8]) rely on adversarial pairs that do not work for HSIC.
- Corollary: Lower bound of $\mathcal{O}\left(\frac{1}{\sqrt{n}}\right)$ on covariance operator estimation, recovering [8].

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