

## Quick Summary

- New non-parametric online change detection algorithm for  $\mathbb{R}^d$ -valued data with runtime complexity of  $O(\log n)$  per observation.
- Guarantees on average run length, uniform false alarm probability, and expected detection delay.
- Minimax optimality of expected detection delay.
- Experimental validation on synthetic, MNIST, HASC, and audio data.
- Key idea: Online approximation of the maximum mean discrepancy on a dyadic grid using random Fourier features.



## Problem Statement

- **Setup:**  $X_1, X_2, \dots, X_t \in \mathbb{R}^d$ ;  $\mathbb{P}, \mathbb{Q}$  probability measures on  $\mathbb{R}^d$ ;  $\mathbb{P} \neq \mathbb{Q}$ .  $\exists \eta \in \mathbb{N} \cup \{\infty\}$  such that

$$X_t \sim \begin{cases} \mathbb{P} & \text{for } t = 1, \dots, \eta \\ \mathbb{Q} & \text{for } t = \eta + 1, \eta + 2, \dots \end{cases}$$

- **Goal:** Stop with minimal delay as soon as  $\eta$  is reached, but not before; never stop in case  $\eta = \infty$ .

## Maximum Mean Discrepancy (MMD)

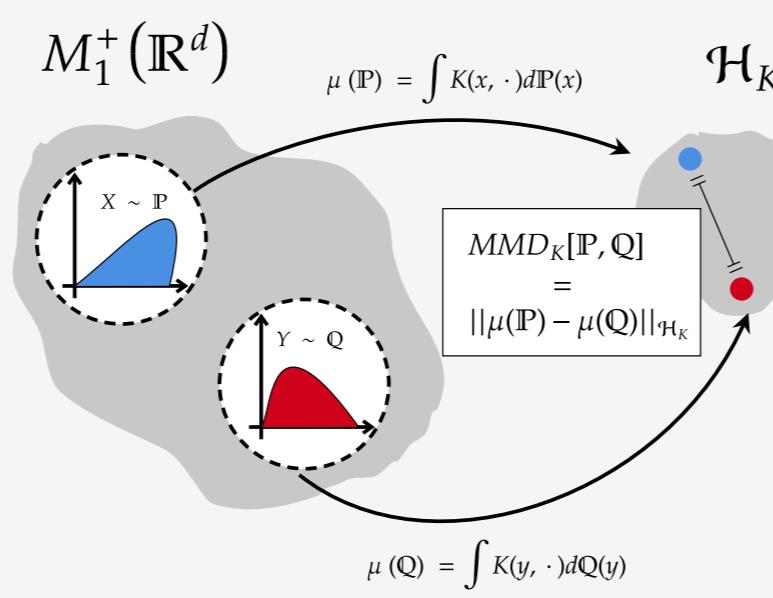
- Let  $K : \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$  be a bounded kernel with associated reproducing kernel Hilbert space (RKHS)  $\mathcal{H}_K$  and (canonical) feature map  $K(\cdot, \mathbf{x})$  for  $\mathbf{x} \in \mathbb{R}^d$ . Then

$$\text{MMD}_K(\mathbb{P}, \mathbb{Q}) = \|\mu_K(\mathbb{P}) - \mu_K(\mathbb{Q})\|_{\mathcal{H}_K},$$

where  $\mu_K : \mathbb{P} \mapsto \int K(\cdot, \mathbf{x}) d\mathbb{P}(\mathbf{x})$  is the kernel mean embedding [1].

- Kernel-based metric on probability measures under mild conditions.

- Classic estimators are  $O(n^2)$ .



## Random Fourier Feature (RFF) Approximation

- If  $K$  is bounded continuous translation-invariant, by Bochner's theorem

$$K(\mathbf{x}, \mathbf{y}) = \int_{\mathbb{R}^d} e^{-i\omega^\top(\mathbf{x}-\mathbf{y})} d\Lambda(\omega).$$

- Approximate  $K$  by [3; 4]

$$\hat{K}(\mathbf{x}, \mathbf{y}) := \langle \hat{z}_K(\mathbf{x}), \hat{z}_K(\mathbf{y}) \rangle, \text{ where } \hat{z}_K(\mathbf{x}) = \frac{1}{\sqrt{r}} ((\sin(\omega_j^\top \mathbf{x}), \cos(\omega_j^\top \mathbf{x})))_{j=1}^r \in \mathbb{R}^{2r},$$

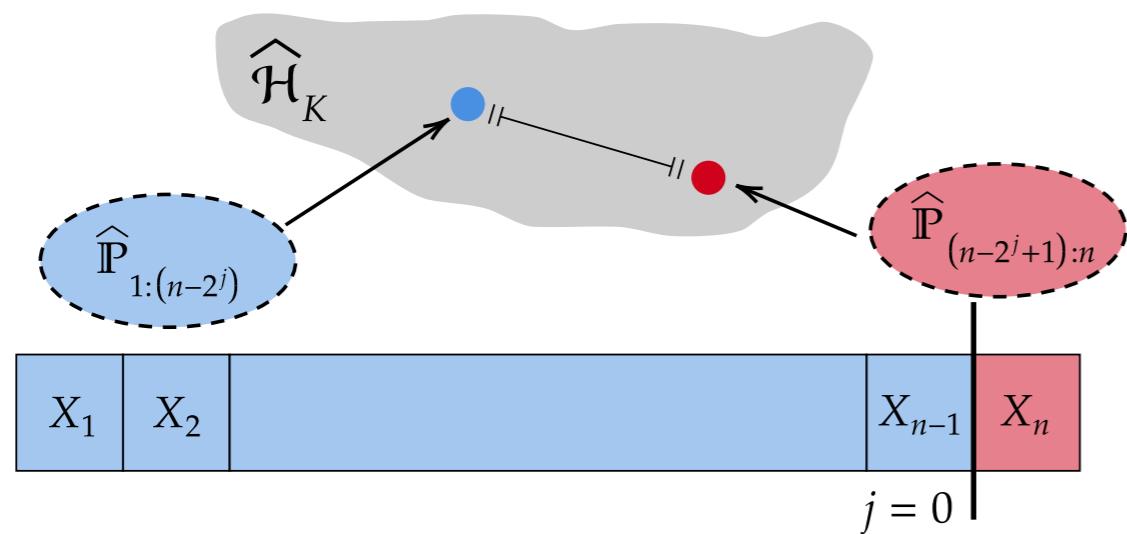
leading to

$$\text{MMD}_{\hat{K}}[X_{1:n}, Y_{1:m}] = \left\| \mu_{\hat{K}}(\hat{\mathbb{P}}_n) - \mu_{\hat{K}}(\hat{\mathbb{Q}}_m) \right\|_{\mathcal{H}_{\hat{K}}} = \left\| \frac{1}{n} \sum_{i=1}^n \hat{z}_K(X_i) - \frac{1}{m} \sum_{i=1}^m \hat{z}_K(Y_i) \right\|_2.$$

- Observation: Both sums can be stored explicitly and MMD can be efficiently computed.

## Proposed Change Detection Algorithm: Online RFF-MMD

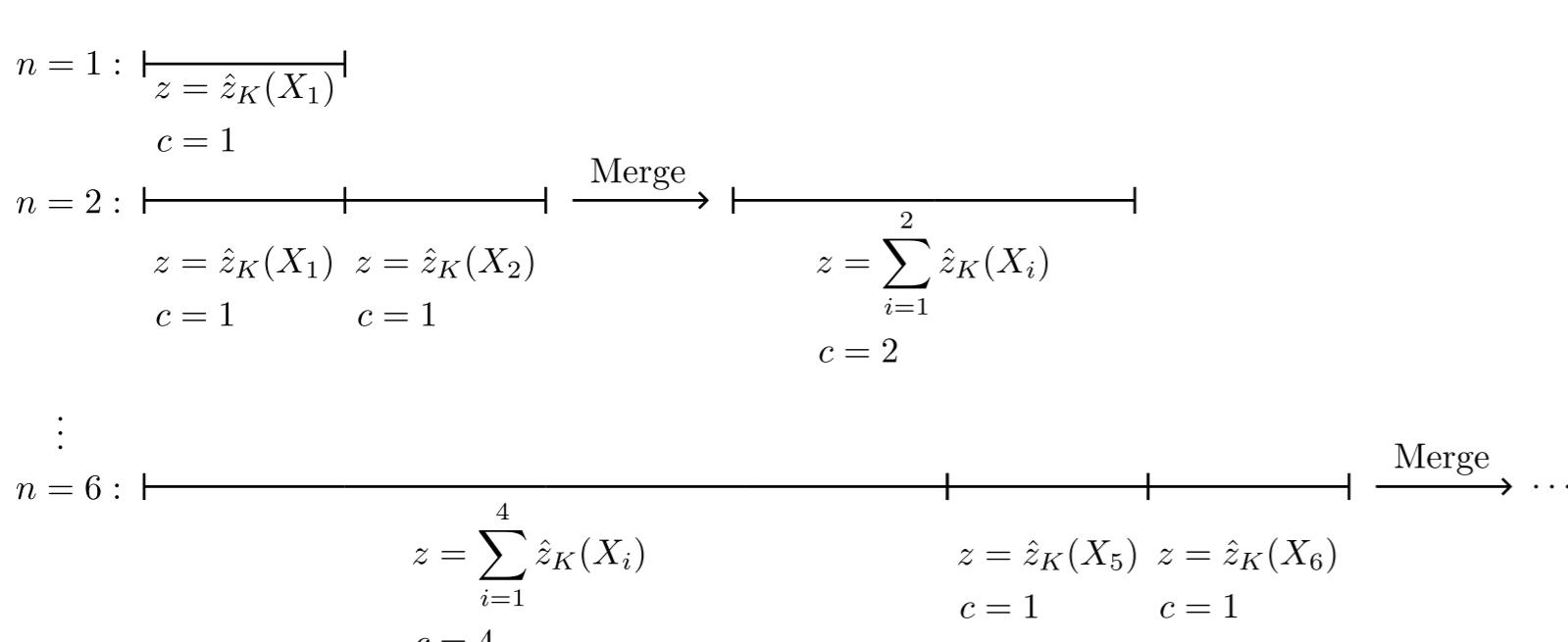
- At the  $n$ -th iteration, the algorithm considers  $\log_2(n)$  splits of the data stream  $\{X_1, \dots, X_n\}$ . For every such split the RFF-MMD between empirical measures of the two samples is computed. The process is stopped at the first  $n$  for which at least one statistic is larger than a given threshold.



- Formally, the Online RFF-MMD stopping time is defined as

$$N = \inf \left\{ n \geq 2 \mid \bigcup_{j=0}^{\lfloor \log_2(n) \rfloor - 1} \sqrt{\frac{2^j(n-2^j)}{n}} \text{MMD}_{\hat{K}}[X_{1:(2^j)}, X_{(2^j+1):n}] > \lambda_n \right\}.$$

- The algorithm has logarithmic time complexity per observation and overall logarithmic space complexity, as illustrated by the following example: upon observing the first  $n = 6$  elements  $X_1, \dots, X_6$ , the algorithm operates as follows



## Theoretical Guarantees

With an appropriately chosen sequence of thresholds, RFF-MMD can be made to attain, respectively, a desired average run length or a desired uniform false alarm probability.

### Average Run Length

For any  $\gamma > 1$ , if the sequence of thresholds satisfies

$$\lambda_n \geq \sqrt{2} + \sqrt{2 \log(4\gamma \log_2(2\gamma))}$$

for all  $n \in \mathbb{N}$ , it holds that  $\mathbb{E}_{\infty}[N] \geq \gamma$ .

### False Alarm Probability

For any  $\alpha \in (0, 1)$ , if the sequence of thresholds satisfies

$$\lambda_n \geq$$

$$\sqrt{2} + \sqrt{2(\log(n/\alpha) + 2 \log \log_2(n) + \log \log_2(2n))}$$

for each  $n \in \mathbb{N}$ , it holds that  $\mathbb{P}_{\infty}(N < \infty) \leq \alpha$ .

With high probability, provided the number of RFFs is chosen sufficiently large, the detection delay is bounded from above by a quantity depending only on the chosen  $\alpha$ , the number of pre-change observations, and the squared MMD between the pre- and post-change distributions.

### Detection Delay

If  $\lambda_n$  is chosen to control the false alarm probability at some level  $\alpha \in (0, 1)$ ,  $\text{supp}(\mathbb{P}) \cup \text{supp}(\mathbb{Q}) \subseteq \mathcal{X}$  for some compact set  $\mathcal{X} \subset \mathbb{R}^d$ , the quantities  $\eta$ ,  $\alpha$ , and  $\text{MMD}_K[\mathbb{P}, \mathbb{Q}]$  jointly satisfy

$$\eta \geq C_1 \frac{\log(2\eta/\alpha)}{(\text{MMD}_K[\mathbb{P}, \mathbb{Q}])^2},$$

and the number of random features is chosen so that

$$\sqrt{r} \geq C_2 \frac{C_3 + \sqrt{2 \log(2/\alpha)}}{(\text{MMD}_K[\mathbb{P}, \mathbb{Q}])^2},$$

then with probability at least  $1 - \alpha$ , it holds that

$$(N - \eta)^+ \leq 1 \vee C_4 \frac{\log(2\eta/\alpha)}{(\text{MMD}_K[\mathbb{P}, \mathbb{Q}])^2},$$

where  $C_1, C_2, C_3$ , and  $C_4$  are absolute constants independent of  $\eta$ ,  $\alpha$ , and  $\text{MMD}_K[\mathbb{P}, \mathbb{Q}]$ .

The detection delay of RFF-MMD is optimal from a minimax perspective, up to logarithmic terms.

## Information Theoretic Bounds

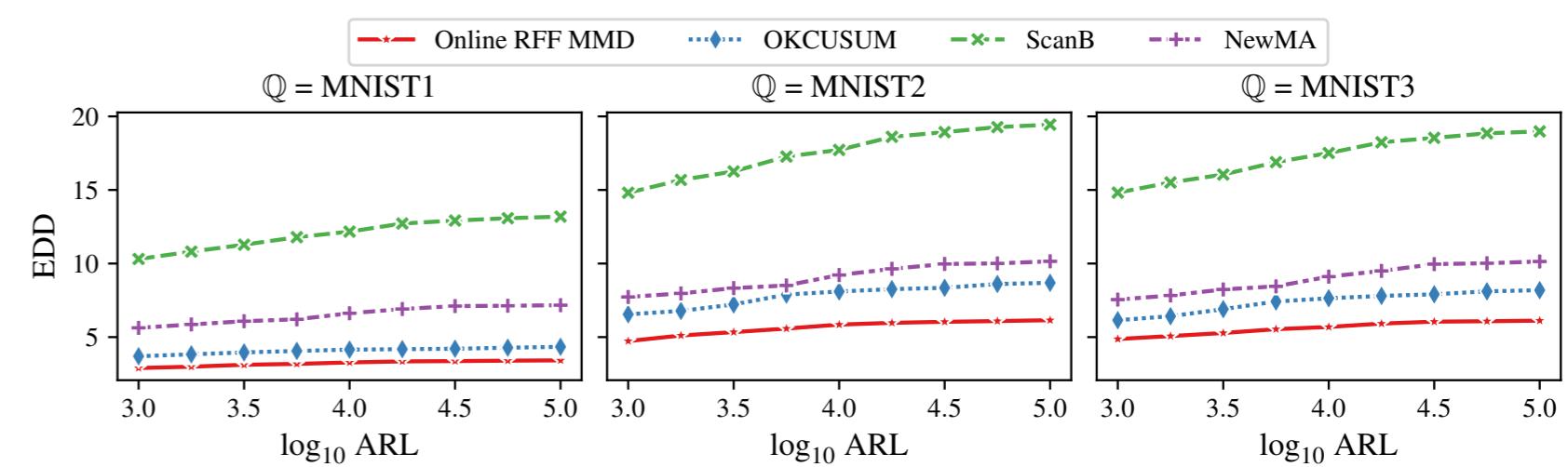
For every bounded, continuous, and translation invariant kernel  $K : \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$  there is a constant  $C_K$  depending only on  $K$  and absolute constants  $\alpha_0, \beta_0 \in (0, 1)$  independent of  $K$ , such that for any  $\alpha \leq \alpha_0$  it holds that

$$\inf_{N: \mathbb{P}_{\infty}(N < \infty) \leq \alpha} \sup_{\substack{\eta > 1 \\ \mathbb{P}, \mathbb{Q} \in \mathcal{M}_1^+}} \mathbb{P}_{\eta} \left( N \geq \eta + C_K \frac{\log(1/\alpha)}{(\text{MMD}_K[\mathbb{P}, \mathbb{Q}])^2} \right) \geq \beta_0$$

with the infimum being over all extended stopping times.

## Numerical Experiments

- **MNIST** ( $d = 768$ ). Pre-change: 64 samples of digit 0; post-change: samples of indicated digit.



- **HASC (Human Activity Sensing Consortium;  $d = 3$ )**. Pre-change: 100 samples of "Walking"; post-change: 100 samples of "Staying".

Algorithm	Average delay	Too early	Miss
Online RFF MMD	<b>21.86</b>	2	1
NewMA	34.25	1	5
ScanB	31.20	0	0
OKCUSUM	17.44	1	0
RuLSIF	20.38	2	0

- **Loudness of Chopin's Mazurka Op. 17 No. 4 ( $d = 1$ )**. Change points in loudness information of Chopin's Mazurkas correspond to score positions having dynamic markings, tempo, or expression markings, among others [2]. Our proposed method flags 10 change points too early, and, on the remaining 15 has an average detection delay of 73.67, with a median detection delay of 64.0.



## References

- [1] A. Gretton, K. Borgwardt, M. Rasch, B. Schölkopf, and A. Smola. A kernel two-sample test. *Journal of Machine Learning Research*, 13(25):723–773, 2012.
- [2] K. Kosta, R. Killick, O. Bandtlow, and E. Chew. Dynamic change points in music audio capture dynamic markings in score. In *International Society for Music Information Retrieval Conference (ISMIR)*, 2017.
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- [4] B. K. Sriperumbudur and Z. Szabó. Optimal rates for random Fourier features. In *Advances in Neural Information Processing Systems (NeurIPS)*, pages 1144–1152, 2015.